

# Patrilocality and Its Effect on Women's Landholding and Household Ownership in India

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## Abstract

This study examines the impact of female marriage migration, a key feature of patrilocal societies, on women's landholding and household ownership rights in India. Despite legislative reforms, such as the Hindu Succession (Amendment) Act of 2005, socio-cultural enforcement of patrilocality continue to annihilate women's property rights. Using district-level household panel data, constructed from multiple datasets, this study employs robust econometric strategies, such as fixed-effects and random-effects panel regression models and regression-based inequality decomposition model, to investigate the association between patrilocality and its relative contribution to perpetuate inequality in women's property ownership across Indian districts. Findings reveal a strong negative association between patrilocality and women's ownership of land and household, reinforcing the "patriarchy-landholding hypothesis and its expansion to other key assets." Robustness checks through sub-sample analyses and endogeneity checks further confirms the validity of the findings. This research fills a gap in understanding the gender disparity in property ownership from a socio-cultural perspective, specifically through the lens of patrilocality.

**Keywords:** Patrilocality, Land Ownership, Household Ownership, Women's Status. India

## Introduction:

In patriarchal societies like India, a woman is often subordinated to the institutions of kinship and marriage, that typically confines them to the private sphere, while men dominate the public regimes. Socially constructed gender-biased roles, often positions men as "providers and bread-winners" and women as "dependents and home-makers" within the family structure, where men hold decision-making authority and control over material resources (Mahato, Das & Reddy, 2025; Garikipati, 2009). These intra-household power dynamics are further entrenched by the unequal inheritance practices that restricts women's access, command, and ownership to ancestral property, due to the continuation of patrilineal descent. Despite, being tethered to two familial affiliations, *i.e.*, her natal family and marital family, she is rarely entitled to the key assets of either household. Deeply entrenched patriarchy favoring kinship and patrilineal inheritance practices, thereby, perpetuate a gendered asset gap, constituting a critical axis of inequality in the development discourse (Agarwal et al., 2021; Rao, 2018).

A central mechanism through which this gendered inequality is sustained in patriarchal societies is '*patrilocality*', a key socio-cultural institution of post-marital residence, where a woman migrates from her natal home to her husband's family home (Agarwal, 1994; 2003). By severing women's physical and emotional proximity to their natal homes/kin, patrilocality not only reinforces patrilineal inheritance claims, but also curtails women's economic ability and social legitimacy to claim over ancestral property. It, thereby, reinforces the normative concept of the male (husband/father-in-law) as the undisputed "head of household", with the land titles and major

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household assets, being overwhelmingly registered under the male head's name, regardless of women's contributions (e.g., [Deere & Doss, 2006](#); [Lastarria-Cornhiel et al., 2014](#)). In this regard, studies also highlight that, a higher perceived socio-economic and emotional utility of sons in providing overall and long-term security to parents, often favours patrilineal inheritance practices and even justifies it by rationalising it through 'dowry practices' as a substitute to their relative natal share, although its translatory value is way lesser ([Agarwal, 1994](#), [Mahato, Das & Reddy, 2025](#)). Particularly in South Asian countries like India, such gendered disparity in property ownership emerges prominently in the ownership of 'land' and 'households', which are often referred as key assets for poor, as they primarily serve as a means to their wealth accumulation ([Gaddis, Lahoti & Swaminathan, 2022](#)).

Estimates from the National Family and Health Survey (NFHS, 2019–21), reveals that not even one-fifth of the total Indian women own land or household in India ([IIPS and MoHFW, 2021](#)). In line with these estimates, recent micro-level estimates by [Mahato, Das & Reddy, 2025](#); underscore/unfolds this gap further, reporting that, compared to 53% and 49% of rural and urban men, only 11% of rural and 15% of urban women, hold any land ownership (agricultural or non-agricultural land ownership) in India. The persistence of such stark disparities, despite gender-sensitive policy interventions and legislative reforms such as the Hindu Succession (Amendment) Act of 2005, granting daughters' coparcenary rights in ancestral property, similar to sons, clearly indicates that legal institutions alone are not sufficient to bridge the gender asset gap, in the absence of institutional support, social recognition, and cultural transformation.

Even when laws entitle women as rightful heirs to ancestral property and joint titles, social expectations and familial dynamics, usually, abstain women from claiming their inheritance rights ([Deninger, Goyal & Nagarajan, 2013](#); [Brule, 2020](#); [Agarwal, 1994](#)). They usually forego their share, due to the fears of estrangement from their natal kin and in particular, their male siblings and the social stigmas attached to it. Limited legal literacy among women, male-dominated bureaucracies and weak enforcement mechanisms to ensure women's de facto ownership over their de-jure ownership, especially in rural areas where socio-cultural forces predominate laws and regulations, further complicate the scenarios ([Deere & Doss, 2006](#)). Additionally, it is also observed that even in cases, where women contribute financially to asset-accumulation, ownership lies with male members. This warrants a need for an in-depth investigation into the social and institutional forces, that continue to act as a structural constraint to women's property ownership.

Given this backdrop, it becomes crucial to examine patrilocality, a key socio-cultural institution of post-marriage residence of women; not merely as a cultural practice, but also as a structural constraint to women's property ownership, as it intersects with inheritance laws, social expectations, and familial obligations, in ways that systematically exclude women from asset accumulation, joint titles and intergenerational wealth transfers. Understanding this linkage is not only crucial from the perspective of gender disparity, wealth accumulation and economic security, but, also in pretext to women's empowerment and development as immovable assets like agricultural land and household ownership are critical to women's agency, dignity and her long-term overall wellbeing.

## **Background and Rationale of the study**

The gendered distribution of property ownership is one of the most contentious debates in the economic development and welfare of women ([Agarwal, 1994](#); [Agarwal, 2003](#); [Brule, 2020](#); [Sen, 1990](#)). Feminist economists and scholars, have eminently contributed to unfold the engrained gender asymmetries in power dynamics, resource allocation and decision-making authority,

initiating within the household, itself that perpetuate gendered gap of critical assets, especially those of economic, social, and political importance. By introducing theoretical framework on non-cooperative bargaining models, and game-theoretic approaches, scholars like [Sen \(1990\)](#) and [Agarwal \(1994, 2003\)](#) dismantles [Gary Becker's New Home Economics \(1971\)](#) and unitary model of household—assuming absolute altruism within the family members and households as a unit of congruent interests, with a maximum single collective utility function and common resource pooling; and reconceptualises households as a site of both co-operation and conflict where intra-household power dynamics and fallback positions, critically shape women's autonomy.

Central to their critique is the idea that access to resources such as land and property are critical to shape women's fallback position, as being asset poor, they are prone to be marginalised and vulnerable, especially in the event of marital discord or family breakdown. As [Agarwal \(2003\)](#) and [Doss et al. \(2014\)](#) argued, land and households are not only a productive asset, but also a critical source of social status, personal security, and negotiating power. In patrilineal and patrilocal systems, where women often lack claim to ancestral land and are excluded from property in their marital homes, this fallback position is effectively nullified and constrained as women are often excluded from property in both their natal and marital homes.

Advancements to a historical-materialist lens, reflected in Friedrich Engels' (1884), *The Origin of the Family, Private Property and the State*, that asserted women's subordination to the rise of private property and patrilineal inheritance, has further been expanded with several theoretical and empirical modification by scholars, proposing a more systemic and theoretically grounded frameworks to analyse patriarchy and women's property ownership. Among those proposed, "landholding-patriarchy" hypothesis, is one of the most prominent and systematic frameworks, exhibiting the linkage between property, kinship, and gendered power, deeply entrenched in patrilineal male predominance. In Indian sub-continent, articulating landholding-patriarchy hypothesis with theoretical insights and empirical assessment, [Goli and Pou \(2014\)](#), showed how concentrated male landownership in rural areas, can diminish women's actual autonomy within households, empirically strengthening Engel's view of land and property as instruments of patriarchal domination, particularly in pre-capitalist and agrarian societies and affirming to Duflo's caution on how formal asset ownership alone doesn't ensure empowerment, unless normative and institutional structures evolve. Although, a number of studies across diverse patrilocal societies such as South Asia and Sub-Saharan Africa have consistently shown that patrilocal, coupled with patrilineal inheritance, is a primary institutional barrier to women's ownership of critical immovable assets in their marital homes ([Agrawal, 1994](#); [Agrawal, 2003](#); [Allendorf, 2007](#); [Lastarria-Cornhiel et al., 2014](#)), scant empirical evidence exists establishing the direct correlation between quantitative measures of "patrilocal" and "women's land ownership rates".

In recent times, however, a growing recognition of patrilocal as a main explanatory variable or a critical measure is observed. For instance, [Khalil and Mookerjee \(2019\)](#), analysed the linkages between patrilocal and women's status and [Licart, \(2023\)](#) analysed it in regards, to women's landholding ownership in Indian context. Unlike previous literature, that were either primary descriptive in nature, or with specific focus, treating patrilocal as a mere background cultural feature, these studies addressed the matter directly. Yet, research gaps persist, as these studies either attempted to provide a comparative picture in regards to other South Asian countries ([Khalil & Mookerjee, 2019](#)) and not referring to property ownership status or examined the linkages at a single time-point and emphasized more on persisting spatial heterogeneity. A major concern, however, lies in their construct measure for patrilocal, that considers patrilocal only under joint family set-up, and using indirect proxy for patrilocal that may not capture patrilocal

effectively. Additionally, they did not account for “female migration due to marriage”, which is a key feature of it. Moreover, these studies fail to capture patrilocality within a nuclear family structure, where sons don’t live with their parents, but married women still move to their husband’s houses, reinforcing the patrilocal culture. This study, thereby, aims to bridge the persisting research gap by quantifying patrilocality more accurately and providing robust empirical evidences for a comprehensive in-depth understanding on women’s ownership of key immovables assets such as landholding and household, under an investigating lens of patrilocality. Following [Goli et al. \(2024\)](#), this study addresses the earlier measurement bias in patrilocality by accounting for marriage induced migration, as a direct measure for patrilocality, which could even capture patrilocality within a nuclear family structure. Employing robust empirical strategies, this study investigates the role of patrilocality in determining women’s ownership status of landholding and households. The objectives of this study stand two- fold:

To investigate the association between patrilocality and women’s ownership status of landholding and households.

To examine the relative contributions of patrilocality in determining the inter-district inequality among women’s ownership in India.

## Methods

### *Data Source:*

Building on the well-established “*landholding-patriarchy hypothesis*” in gender studies, we investigate the association between patrilocality and women’s property ownership, emphasizing on two key immovable assets of economic relevance: land and household, in India. For an empirical assessment, patrilocality, the key predictor, has been estimated utilizing information on the migrants by their respective last place of residence and the reason stated for migration, classified by age and sex; from the D-05 series of the Census of India (2001 & 2011).

Its operational definition and the measure of construct has been discussed in the later section. Further, the information on key outcome variables, i.e., land and house ownership alongside its correlates, including individual and household characteristics, have been drawn from the last two conducted surveys of the National Family and Health Surveys (NFHS) for the survey years, 2015-16 and 2019-21. Integrating this information, a district level panel dataset of 619 districts have been constructed for this study. Table 1 provides a detailed description of variables and its data source for this study.

*Predictor Variable:* Following [Goli et al., 2024](#); ‘patrilocality’ is constructed as the ratio of female marriage migrants and male marriage migrants aged 10 to 49 years, at the district level in India. Mathematically it can be expressed as,

$$Patrilocality = \left( \frac{Female\ Marriage\ Migration_{10-49yrs.}}{Male\ Marriage\ Migration_{10-49yrs.}} \right)$$

where, marriage migration refers to the number of females and males in the age-group of 10 years to 49 years, respectively, who migrated from the place of their last residence owing to marriage. This could be—within the state of enumeration, outside the district of enumeration within the state, and elsewhere within the district of enumeration.

### *Outcome Variables:*

*Landholding Ownership:* Using household questionnaire, the information on key outcome variables of this study, *i.e.*, landholding ownership has been extracted, based on the information available on “Who owns this agricultural land?” The information has been collected under the categories as follows: Male member, Female Member, Both and Don’t Know. On clubbing together the categories, “female member” and “both”, women’s landholding ownership was estimated, if they owned it independently or jointly.

*Household Ownership:* Similarly, women’s household ownership was estimated, if they owned the house independently or jointly, based on the information from the household questionnaire on “Who owns this house?”, classified under the categories as Male member, Female Member, Both and Don’t Know. However, we acknowledge the fact that some studies have argued: the very concept of "household ownership" is problematic in patrilocal contexts. It masks intra-household power imbalances and renders women's economic contributions and rights invisible. Research shows that assets controlled solely by women within the household are often minimal compared to those controlled by men or jointly (Quisumbing & Maluccio, 2003).

*Correlates:* Based on the extensive literature review on gender asset gap, key socio-demographic, individual and household characteristics has been used to capture the true essence of the linkages between patrilocality and women’s landholding as well as household ownership (see Table 1).

### *Empirical Strategy*

Utilizing the district level panel dataset, this study employs several statistical tools to empirically examine the role of patrilocality in women’s property (land & household) ownership. Initiating at first, with the summary statistics table, we provide the overview of the variables used in the study and have performed paired samples t-test statistics to observe the within-unit changes over time for two distinct time-points and assess, if the observed change is statistically significant. Statistical models employed for empirical assessment and its specifications have been mentioned below.

### *Fixed Effects and Random Effects Panel Data Regression*

Unadjusted and adjusted fixed-effects (FE) and random-effects (RE) panel data regression to examine the linkages between patrilocality and women’s ownership for land and household, separately. Although the Hausman test, performed to choose between FE and RE models, suggest for FE effects, we have also performed RE panel regression models, as it allows for the inclusion of both time-variant and time-invariant variables, that are essential for understanding structural and cultural determinants of women’s property ownership in India. As this study uses limited time periods with just two time points, Hausman test may be sensitive to issues such as heteroskedasticity. Thus, to ensure precise estimates and better efficiency of the models to capture minor coefficient differences, RE model’s estimates have been reported, alongside FE models. General equations for the same is as mentioned:

Fixed-Effects models (adjusted):

$$Land\_Ownership_{it} = \beta_0 + \beta_1 Patrilocality_{it} + \beta_2 Control\_variables_{it} + \beta_3 Time\_dummy + \lambda_t + (u_i + \varepsilon_{it})$$

$$House\_Ownership_{it} = \beta_0 + \beta_1 Patrilocality_{it} + \beta_2 Control\_variables_{it} + \beta_3 Time\_dummy + \lambda_t + (u_i + \varepsilon_{it})$$

Where, Landholding Ownership<sub>it</sub> and Household Ownership<sub>it</sub> represents the women’s landholding and household ownership status, respectively for district *i* for the time-period *t*.  $\lambda_t$  captures time-fixed effects.  $u_i$  are entity-specific fixed effects at the district level for all time-invariant characteristics and  $\varepsilon_{it}$  are the standard errors clustered by entity.

Random Effects models (adjusted):

$$Land\_Ownership_{it} = \beta_0 + \beta_1 Patrilocality_{it} + \beta_2 Control\_variables_{it} + \beta_3 Time\_dummy + (u_i + \varepsilon_{it})$$

$$House\_Ownership_{it} = \beta_0 + \beta_1 Patrilocality_{it} + \beta_2 Control\_variables_{it} + \beta_3 Time\_dummy + (u_i + \varepsilon_{it}),$$

$$and\ u_i \sim N(0, \sigma_u^2)$$

Where, Landholding Ownership<sub>it</sub> and Household Ownership<sub>it</sub> represents the women's landholding and household ownership status, respectively for district *i* in period *t*, respectively.  $\beta$  is the coefficient for independent variables.  $u_i$  ( $i=1 \dots n$ ) is a random individual-specific effect at the district level, uncorrelated with the regressors and  $\varepsilon_{it}$  is the error term for respective time period.

### *Regression based Inequality Decomposition Model*

Next, to meet our second objective, we employ Regression-based Inequality Decomposition model, proposed by Shorrocks (1982) and further extended by Fields (2003), to understand the relative contribution of patrilocality in determining inter-district inequality in women's property ownership, in regards to land and house. The detailed equations for the same is as follows:

Mathematically,

$$\widehat{\mu}_{ij} = \alpha + \beta_{1ij} * x_{1ij} + \beta_{2ij} * x_{2ij} + \varepsilon$$

where  $\widehat{\mu}_{ij}$  denotes predicted values for women's landholding and household ownership, respectively, as a function of its components, including patrilocality and  $\varepsilon$  is the residual term. The relative contribution of the patrilocality and other correlates components can be written as –

$$\sigma^2(\widehat{\mu}_{ij}) = \sum_{i=1}^k \beta_i \text{cov}(\widehat{\mu}_{ij}, x_i) + \sigma^2(\varepsilon)$$

Here,  $\sigma^2(s)$  is the variance of  $\widehat{\mu}_{ij}$ ,  $\text{cov}(\widehat{\mu}_{ij}, x_i)$  is the covariance of  $\widehat{\mu}_{ij}$  with each variable  $x_i$ , for which the relative contributions cumulate to 100 percent, on accounting for the residuals. This decomposition function is analyzed using the *ineqrbd* package of the STATA software, *version 19.5*.

## **Results**

### ***Descriptive Findings:***

Summary statistics provided in Table 1, suggest an observable changes over time for the time-period 2001 and 2011, for the variables used in this study<sup>3</sup>. Patrilocality, the predictor variable, defined as the practice of married couples living with the husband declined sharply (-45.23\*\*\*) family—declined sharply (-45.23), signaling a shift away from traditional household structures. Meanwhile, key outcomes, i.e., women's landholding and household ownership saw modest increases (0.88 and 1.14, respectively), suggesting small improvements in property ownership over time, though the change being gradual. The statistically insignificant change for the land ownership could plausibly be understood from the structural rigidity in land ownership patterns owing legal complexity, bureaucratic and social hinderances, favoring male lineage in inheritance patterns, leading to the delayed improvement in ownership despite amendments in land reforms. Additionally, it may have also occurred, owing to the high heterogeneity across the districts, thereby leading to statistically insignificant estimate due to very small net national average change for landholding ownership among women. Besides, several other correlates have also undergone significant changes over time, reflecting evolving social norms and family dynamics, modest

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<sup>3</sup> The means provided in Table1 are aggregated at district level and the averages may not be nationally representable.



economic and development progress and shifting demographic patterns, over time. Next, we examined the women's relative ownership status of landholding or households, based on the level of patrilocality, prevailing across the districts in India. From mosaic plots shown in Figure 1, it is clearly observed that women's ownership status decreases with rise in patrilocality as the proportionate share of women falling under higher ownership status for both land and household is higher across the districts with relatively low levels of patrilocality and vice versa. This sets the premise for an empirical assessment of a) the linkages between patrilocality and women's property ownership status, suggesting an inverse relationship between the two and b) examining inter-district inequality in women's property ownership, as ownership status for women across districts, varied by patrilocality level.

### ***Main Findings:***

#### *Patrilocality and Women's Ownership of Landholding and Household in India*

Tables 2 and 3 provide unadjusted and adjusted FE and RE panel data regression model estimates, examining the association between patrilocality and women's ownership of Landholding and Household across districts in India. The estimates clearly suggest a negative and statistically significant association between patrilocality and women's landholding ownership across all models, despite controlling for key controls. It suggests that a rise in one unit change for patrilocality can decrease the likelihood of women owning land by around 1.2% to 2% across the districts in India. Similarly, for household ownership as well, the negative and statistically significant association, holds true except for last FE model, adjusted for time-fixed entity. Estimates suggest that one unit rise in patrilocality can decrease the likelihood of women owning a household by around 0.6% to 1.3%. The findings, thereby clearly strengthen our argument that patrilocality acts as a structural barrier in women's property ownership rights by reducing their probability to own land or households, thereby, warranting a need for context-specific tailored interventions to raise legal and community awareness regarding women's property ownership rights, in regards to land and house, the two critical assets for economic resource, especially for asset poor household (Agarwal, 2021).

#### *Inter-district inequality in Women's Ownership of Landholding and Household in India*

Figures 2 and 3 shows the inter-district inequality and relative contribution of the key factors attributing to this inequality, in regards to women's ownership of landholding and household in India. As observed, moderate level of inequality is observed for both household and land ownership among women across the district (Figure 2), with inequalities, slightly being higher in landholding ownership, compared to household ownership across the districts. Among all explanatory variables, patrilocality emerged as the most significant determinant with largest individual contribution of the explained variance for inter-district inequality, attributable to women's landhold (28%) and household (24%) ownership, respectively. Other notable contributors to this inequality, include mean household size, educational status and mean age at marriage (Tables 4 and 5). Substantial inequality, explained by residual in both cases may have arisen due to unobserved cultural, institutional and legal factors, besides non-inclusion of policy-related interventions to improve women's property ownership rights and their effective access and control. Moreover, relatively lesser contribution of socio-economic and development factors such as wealth status and residing in urban households suggest that improvements in socio-development factors without addressing the deeply entrenched patriarchal hierarchical set-up of cultural institutions like patrilocality, wouldn't yield desired improvements in women's property ownership.

### Robustness Checks:

Robustness Checks through heterogeneity analyses examining the association between patrilocality and women's property ownership of land as well as household by region (Figures 3, 4 and 5), further confirm the robustness and validity of our main findings presented in Tables 2, 3 and 5; highlighting the deterring influence of patrilocality on women's ownership status of land and household. As observed across both wheat and rice region, the association of patrilocality with women's ownership of both household and landhold is found to be statistically significant and negative (Figures 3).

Thus, in alignment with the theoretical works of feminist economics and anthropological studies on property inheritance in India, highlighting the critical role of patriarchy and pertinent regional variations (Agarwal, 1994; Kelkar & Nathan, 1991), our study contends that patrilocality serves as one of the most prominent institutional determinants of gendered asset gap and its variations across regions in India.

### Conclusions

This study contends that building a strong women's agency with effective implementation of gender-equitable property rights, for key immovable assets such as "landholding" and "household ownership", efforts shall move beyond addressing individual level factors and concerned interventions must consider for the structural constraints imposed by social-cultural norms, such as patrilocality. On recognizing, patrilocality, as a key structural constraint and estimating it through a direct measure of marriage-induced migration for women following Goli et al., 2024, we highlight the need for policy interventions that address not only legal entitlements but also residential patterns, mobility constraints, and the gendered geography of land markets. Additionally, findings call for more targeted and tailored policy programs, especially for regions with stronger adherence to patrilocal norms, where women may face restrictions to social mobility lack social support and connects to local institutions. The study, thereby, contributes significantly to the literature on gendered asset gap in India, both conceptually and empirically, on offering a more robust measure for capturing the true essence of patrilocality and centralising its role as a predominant socio-cultural institution in perpetuating inter-gender as well as intra-gender asset gap in property ownership.

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### Appendix: Tables and Figures

Table 1. Summary Statistics of the Study Variables Used in this Study

Variable	2001		2011		Difference in Means	Data Source
	Mean	Std. Dev.	Mean	Std. Dev.		
Patrilocality	101.59	67.47	55.74	31.21	45.23***	Census of India: (2001) & (2011)
Landholding Ownership	11.34	7.38	12.07	7.07	-0.88	NFHS: (2015-16), (2019-21)
Household Ownership	15.19	7.64	16.13	8.28	1.14*	NFHS: (2015-16), (2019-21)
Mean Household Size	4.75	0.61	4.51	0.64	0.27***	NFHS: (2015-16), (2019-21)
Nuclear Households	57.29	6.66	57.61	7.40	-0.62	NFHS: (2015-16), (2019-21)
Mean Years of Schooling(>=10yrs)	35.14	14.26	40.95	14.20	-6.31***	NFHS: (2015-16), (2019-21)
Low Wealth Quintile Households	32.56	23.68	30.62	22.25	1.00	NFHS: (2015-16), (2019-21)
Mean Age at Marriage	18.31	1.22	18.67	1.29	-0.44***	NFHS: (2015-16), (2019-21)
Urban Households	27.82	22.62	31.16	21.64	-2.84***	Census of India (2001), (2011)
Son Preference	19.15	9.92	15.49	8.28	3.38***	NFHS: (2015-16), (2019-21)
Hindu	80.17	17.96	81.14	17.72	-0.49	NFHS: (2015-16), (2019-21)
Muslims	13.79	14.50	13.35	14.49	0.49	NFHS: (2015-16), (2019-21)
SCs	21.12	9.76	22.93	9.83	-1.61**	NFHS: (2015-16), (2019-21)
STs	10.15	15.95	10.08	15.71	0.04	NFHS: (2015-16), (2019-21)

Note: \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Figure 1. Mosaic Plots showing the association between Patrilocality and Household as well as Women's Landholding Ownership in India

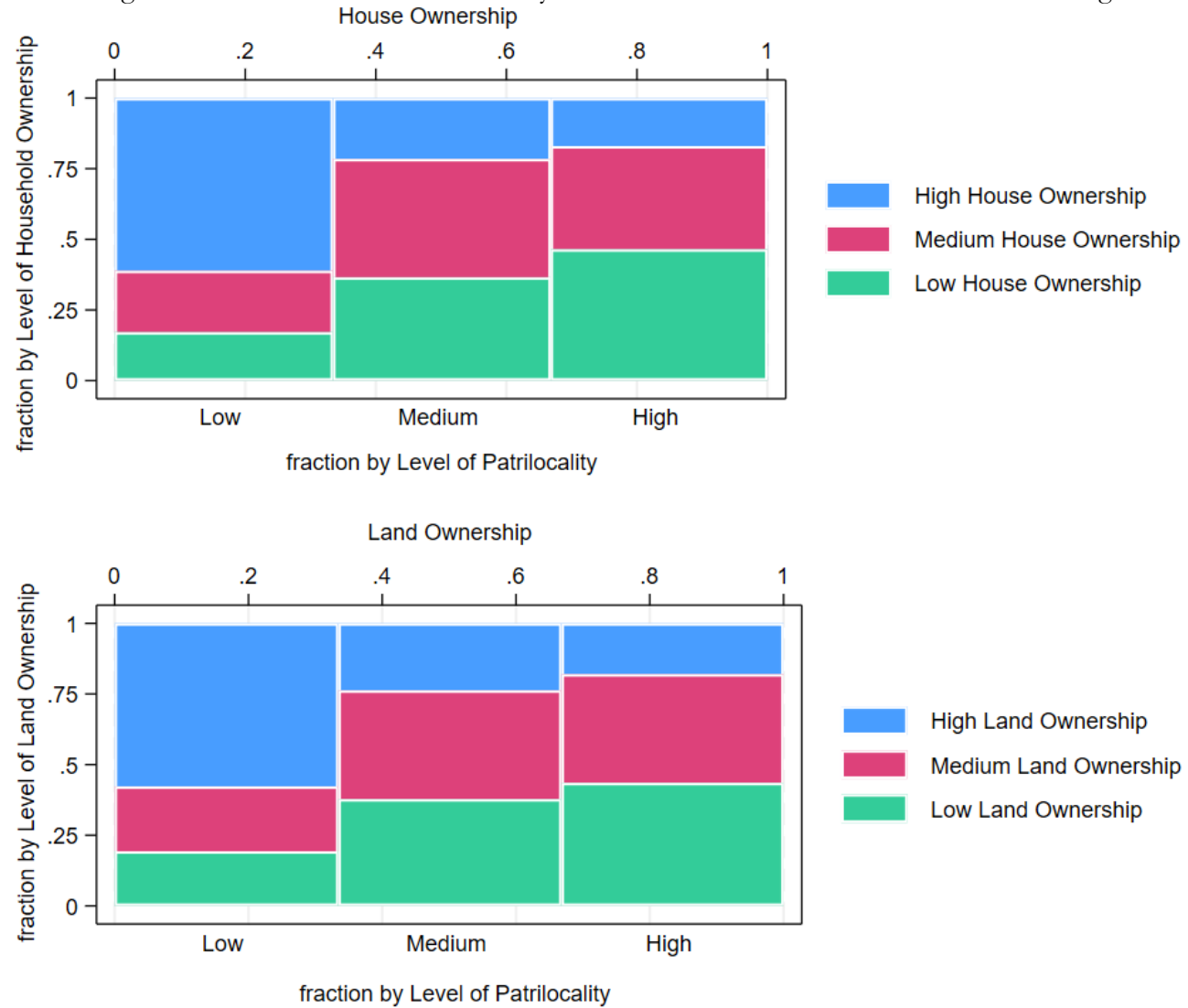


Table 2. Fixed Effects & Random-Effects Panel Regression Model Estimates, assessing the relationship between Patrilocality and Land Ownership of Women in India

Variables	Fixed Effects			Random Effects		
	Baseline	Adjusted <sup>1</sup>	Adjusted <sup>2</sup>	Baseline	Adjusted <sup>1</sup>	Adjusted <sup>2</sup>
Patrilocality	-0.0150*** (0.00281)	-0.0120*** (0.00351)	-0.00824** (0.00413)	-0.0202*** (0.00262)	-0.0145*** (0.00311)	-0.0136*** (0.00365)
Son Preference			0.0244 (0.0348)		-0.0727** (0.0302)	-0.0844*** (0.0308)
Mean Household Size		-1.663** (0.792)	-1.081 (0.886)		0.876 (0.656)	1.418** (0.703)
Nuclear Households		-0.109** (0.0492)	-0.0887* (0.0519)		0.0925** (0.0430)	0.0944** (0.0437)
Mean Years of Schooling ( $\geq 10$ yrs.)		0.0454 (0.0396)	0.00858 (0.0451)		0.147*** (0.0302)	0.154*** (0.0318)
Households under Low Wealth Quintile		0.0265 (0.0227)	0.0125 (0.0242)		-0.0235 (0.0173)	-0.0289 (0.0186)
Mean Age at Marriage		-1.008** (0.423)	-1.100** (0.426)		-1.350*** (0.287)	-1.240*** (0.289)
Urban Households		-0.0215 (0.0369)	-0.0142 (0.0371)		-0.0137 (0.0187)	-0.0182 (0.0188)
Hindus		0.0293 (0.0491)	0.0360 (0.0493)		-0.0955*** (0.0176)	-0.0833*** (0.0179)
Muslims		0.0659 (0.0628)	0.0619 (0.0628)		-0.107*** (0.0234)	-0.101*** (0.0234)
SCs		0.0126 (0.0291)	0.00310 (0.0299)		-0.0672*** (0.0251)	-0.0682*** (0.0253)
STs		0.00711 (0.0345)	0.00514 (0.0347)		0.0443** (0.0177)	0.0453** (0.0176)
Wheat Belt						-2.363*** (0.750)
Time-Fixed Entity			0.794* (0.449)			-0.00692 (0.0335)
Constant	13.37*** (0.240)	40.73*** (11.25)	38.86*** (11.32)	13.78*** (0.423)	35.24*** (7.896)	44.88 (68.12)
Observations	1,238	1,238	1,238	1,238	1,238	1,238
R-squared	0.044	0.065	0.070			
Number of Id	619	619	619	619	619	619

Standard errors in parentheses: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table 3. Fixed Effects &amp; Random-Effects Panel Regression Model Estimates assessing the relationship b Patrilocality and Household Ownership status of Women in India

Variables	Fixed Effects			Random Effects		
	Baseline	Adjusted <sup>1</sup>	Adjusted <sup>2</sup>	Baseline	Adjusted <sup>1</sup>	Adjusted <sup>2</sup>
Patrilocality	-0.0136*** (0.00261)	-0.00760** (0.00315)	-0.000280 (0.00365)	-0.0194*** (0.00249)	-0.0105*** (0.00292)	-0.00686** (0.00341)
Son Preference			-0.0415 (0.0308)		-0.118*** (0.0283)	-0.130*** (0.0288)
Mean Household Size		-4.596*** (0.709)	-3.082*** (0.784)		-2.025*** (0.628)	-0.793 (0.667)
Nuclear Households		-0.290*** (0.0440)	-0.226*** (0.0459)		-0.0524 (0.0408)	-0.0345 (0.0412)
Mean Years of Schooling (>=10 yrs.)		-0.000988 (0.0355)	-0.0763* (0.0399)		0.0961*** (0.0292)	0.0965*** (0.0308)
Households with Low Wealth Quintile		0.0493** (0.0203)	0.0343 (0.0214)		-0.00809 (0.0166)	-0.0281 (0.0178)
Mean Age at Marriage		-0.973** (0.379)	-1.163*** (0.377)		-1.590*** (0.282)	-1.429*** (0.280)
Urban Households		0.0358 (0.0330)	0.0496 (0.0328)		0.0442** (0.0188)	0.0386** (0.0186)
Hindus		-0.0339 (0.0440)	-0.0167 (0.0436)		-0.103*** (0.0178)	-0.0788*** (0.0177)
Muslims		0.0247 (0.0562)	0.0247 (0.0556)		-0.104*** (0.0238)	-0.0933*** (0.0234)
SCs		0.00173 (0.0261)	-0.0253 (0.0265)		-0.0513** (0.0238)	-0.0568** (0.0238)
STs		0.0124 (0.0309)	9.04e-05 (0.0307)		0.0252 (0.0175)	0.0262 (0.0171)
Wheat Belt						-5.075*** (0.759)
Time-Fixed Entity			1.420*** (0.397)			0.0295 (0.0317)
Constant	16.29*** (0.224)	71.59*** (10.08)	66.36*** (10.02)	16.75*** (0.432)	65.56*** (7.692)	-1.866 (64.60)
Observations	1,238	1,238	1,238	1,238	1,238	1,238
R-squared	0.042	0.131	0.157			
Number of Id	619	619	619	619	619	619

Standard errors in parentheses: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Figure 2: Inequality in Women's House and Land Ownership across Indian Districts: Lorenz Curves and Gini Coefficients

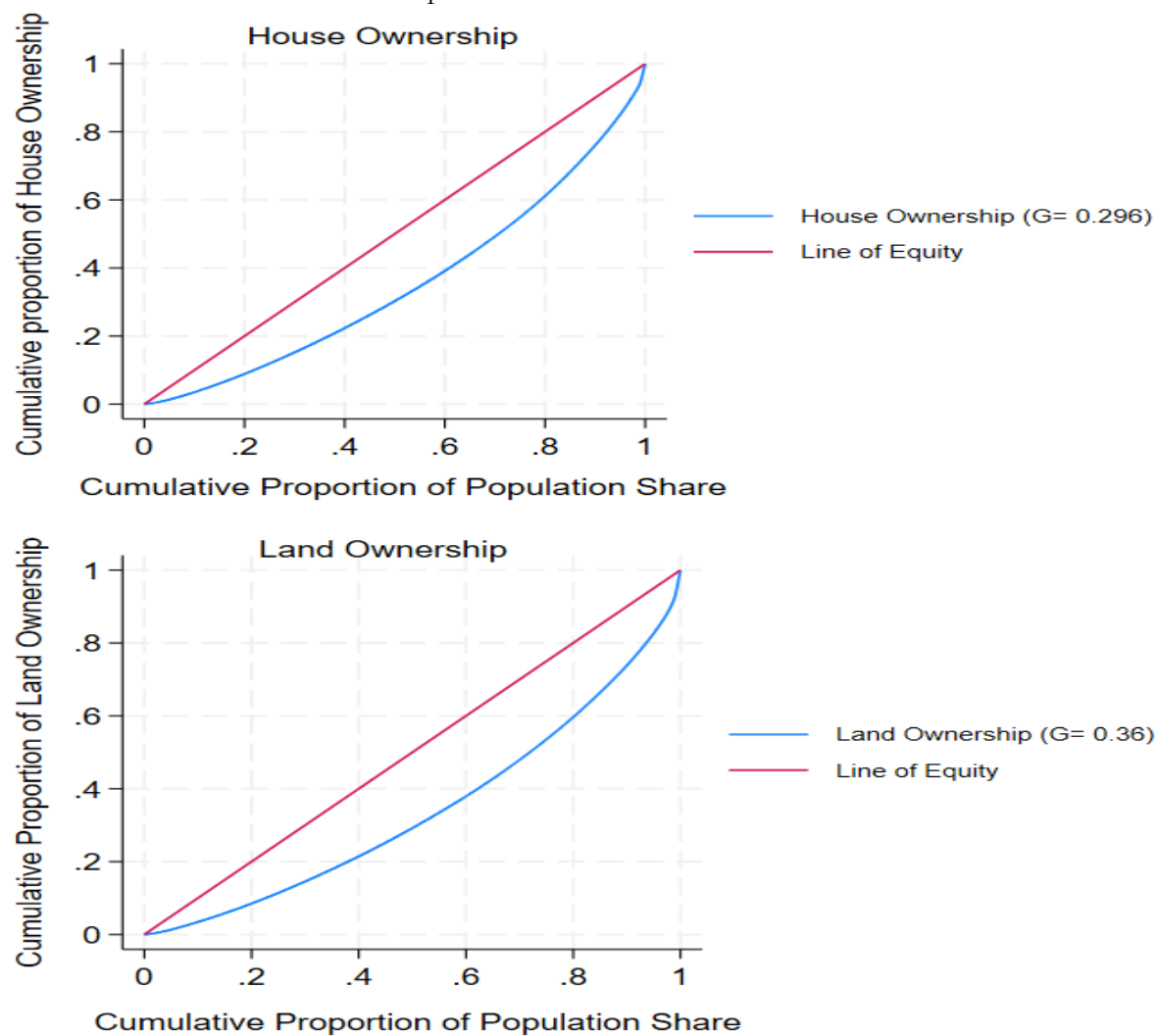




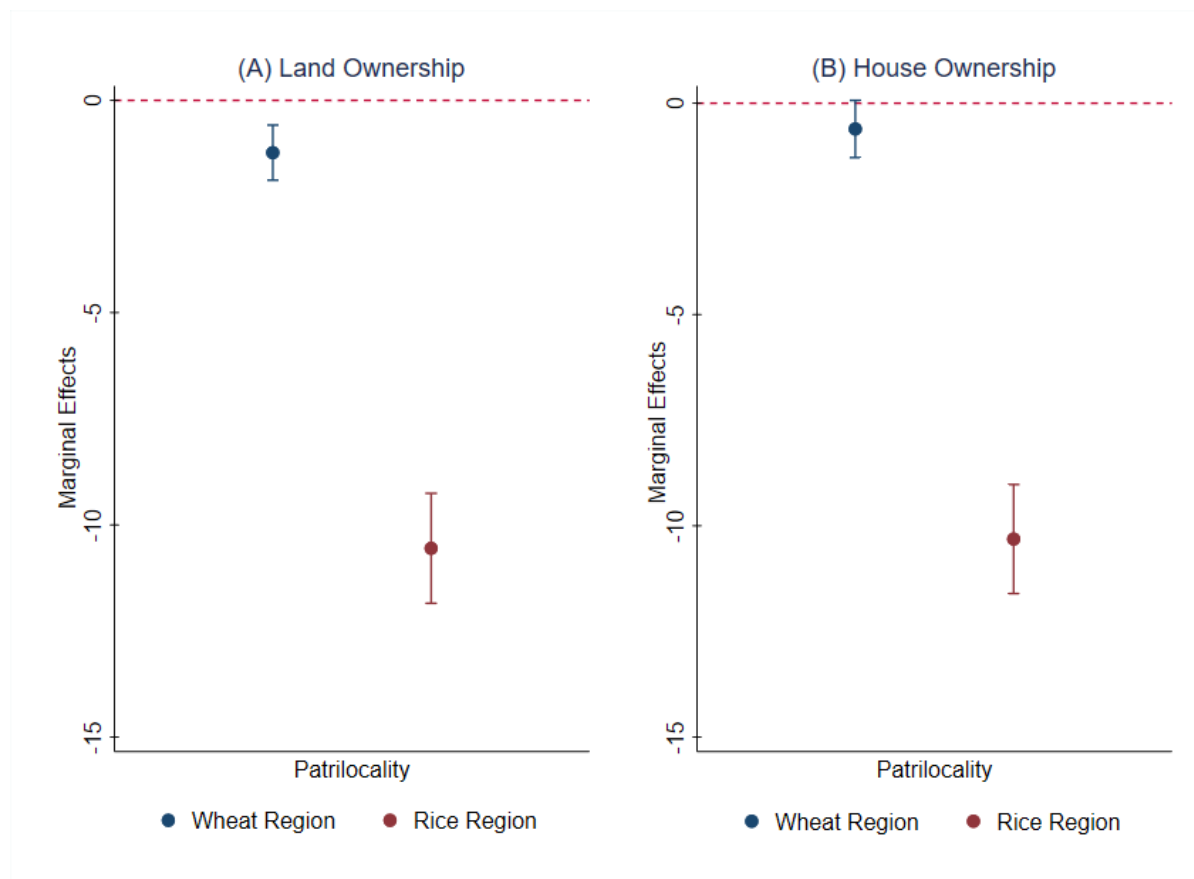
Table 4. Regression Based Inequality Decomposition Estimates: Relative Contribution of Patrilocality and Other factors to inequality in Women's Landholding and Household Ownership in India

	Landholding Ownership	Household Ownership
	Relative Percentage Contribution	Relative Percentage Contribution
Patrilocality	28.16*** (0.2983)	24.01*** (0.2972)
Son Preference	2.90*** (0.0327)	4.29*** (0.0327)
Age at Marriage	3.59*** (0.2279)	3.56*** (0.2279)
Mean Household Size	5.91*** (0.6496)	6.87*** (0.6472)
Family Structure	3.37*** (0.0429)	3.46*** (0.0427)
Educational Status	5.36*** (0.026)	6.23*** (0.0259)
Wealth Status	0.67 (0.0161)	2.36*** (0.016)
Urbanisation	0.93*** (0.2288)	1.30*** (0.0134)
Religion	5.07*** (0.0134)	3.03 (0.0132)
Caste	5.50* (0.0252)	2.43*** (0.0251)
Region	2.70*** (0.1572)	6.68*** (0.0143)
Time	0.39*** (0.0439)	0.47*** (0.0437)
Residual	35.44	35.32
Number of observations	1,238	1,238
F (14, 1223)	95.42	113.49
Prob > F	0	0
R-squared	0.5221	0.5651
Adjusted R-squared	0.5166	0.5601

Standard errors in parentheses: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

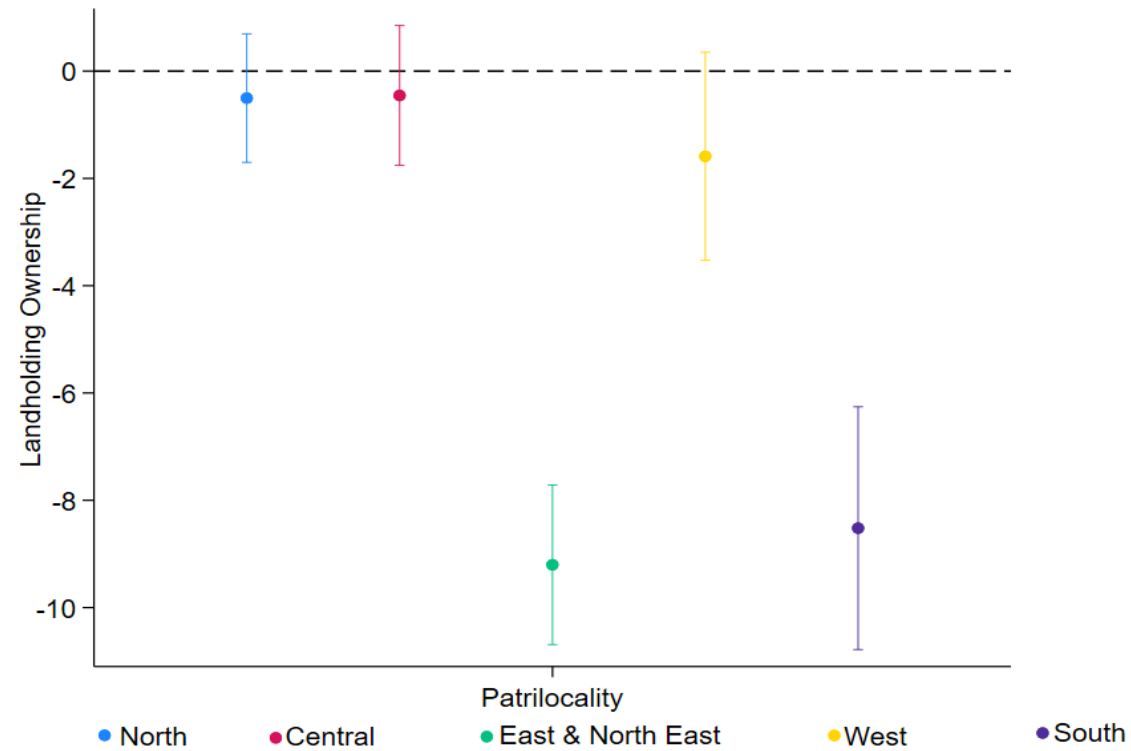
### Heterogeneity Analyses:

Figure 3. Marginal Effects Regression Plots showing the association between Patrilocality and Women's Ownership of Land and House by Wheat and Rice Belt Regions



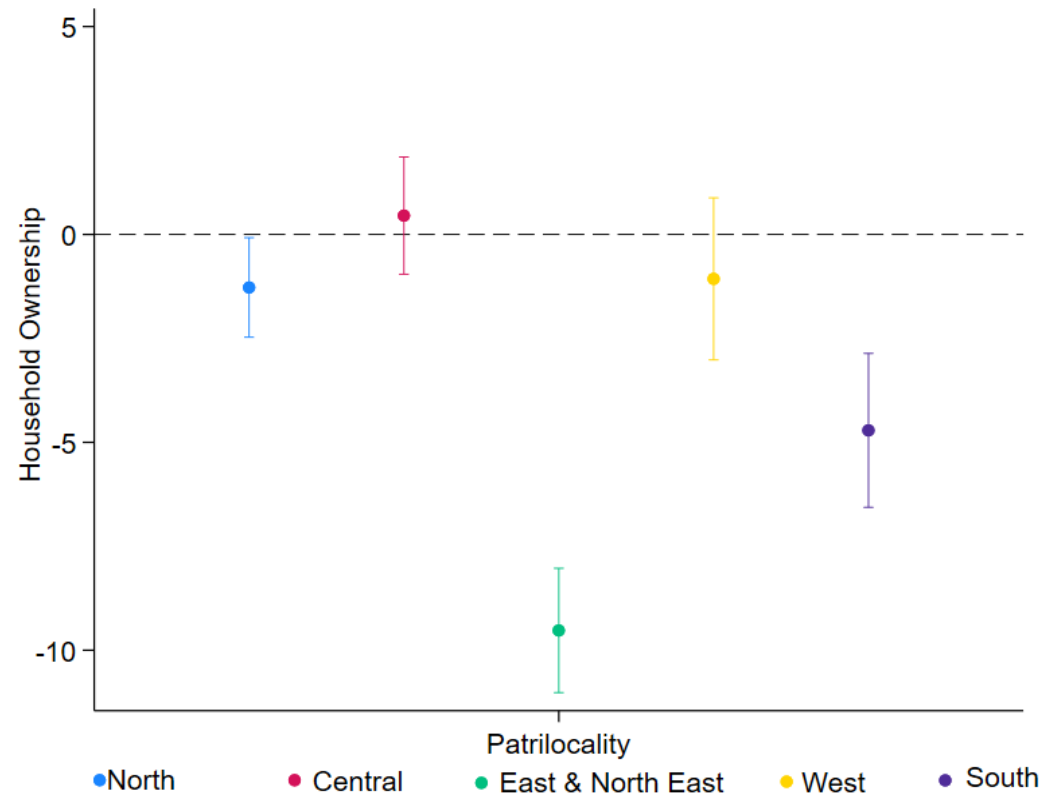
*Note:* The model is adjusted for all key controls considered in Table 2 and 3, including time

Figure 4. Co-efficient Plots showing the relationship between Patrilocality and Women's Landholding Ownership across major regions in India



*Note:* The model is adjusted for all key controls considered in Table 2 and 3, including time

Figure 5. Co-efficient Plots showing the relationship between Patrilocality and Women's Household Ownership across major regions in India



*Note:* The model is adjusted for all key controls considered in Table 2 and 3, including time